LARGE-SAMPLE THEORY

The distribution of a function of several sample means, e.g.

$$g(\bar{X}, \bar{Y})$$

is usually too complicated. The central limit theorem states that this distribution tends, as $N \to \infty$, to a Normal distribution with the mean of

$$g(\mu_x, \mu_y) + \dots$$

and a variance given by

$$\begin{split} &\left(\frac{\partial g(\mu_x,\mu_y)}{\partial \mu_x}\right)^2 \mathrm{Var}(\bar{X}) + \left(\frac{\partial g(\mu_x,\mu_y)}{\partial \mu_y}\right)^2 \mathrm{Var}(\bar{Y}) \\ &+ 2\frac{\partial g(\mu_x,\mu_y)}{\partial \mu_x} \frac{\partial g(\mu_x,\mu_y)}{\partial \mu_y} \mathrm{Cov}(\bar{X},\bar{Y}) + \dots \end{split}$$

When dealing with a random independent sample of X, Y pairs (a special case which does NOT apply to to time-series formulas quoted below), we get

$$\left(\frac{\partial g(\mu_x,\mu_y)}{\partial \mu_x}\right)^2 \frac{\sigma_x^2}{N} + \left(\frac{\partial g(\mu_x,\mu_y)}{\partial \mu_y}\right)^2 \frac{\sigma_y^2}{N} + \frac{\partial g(\mu_x,\mu_y)}{\partial \mu_x} \frac{\partial g(\mu_x,\mu_y)}{\partial \mu_y} \frac{\rho_{xy}\sigma_x\sigma_y}{N} + \dots$$

The general formula can be derived from

$$\begin{split} g(\bar{X},\bar{Y}) &\simeq g(\mu_x,\mu_y) + \frac{\partial g(\mu_x,\mu_y)}{\partial \mu_x} \left(\bar{X} - \mu_x\right) + \frac{\partial g(\mu_x,\mu_y)}{\partial \mu_y} \left(\bar{Y} - \mu_y\right) \\ &+ \frac{1}{2} \frac{\partial^2 g(\mu_x,\mu_y)}{\partial \mu_x^2} \left(\bar{X} - \mu_x\right)^2 + \frac{1}{2} \frac{\partial^2 g(\mu_x,\mu_y)}{\partial \mu_y^2} \left(\bar{Y} - \mu_y\right)^2 \\ &+ \frac{\partial^2 g(\mu_x,\mu_y)}{\partial \mu_x \partial \mu_y} \left(\bar{X} - \mu_x\right) \left(\bar{Y} - \mu_y\right) + \dots \end{split}$$

by realizing that

$$\begin{array}{rcl} \mathbb{E}\left(\bar{X}-\mu_x\right) & = & 0 \\ \\ \mathbb{E}\left[\left(\bar{X}-\mu_x\right)^k\right] & = & O\left(\frac{1}{N^{\left[\frac{k+1}{2}\right]}}\right) \end{array}$$

The error of this approximation is of the $O\left(\frac{1}{\sqrt{N}}\right)$ type. To remove this error, on has to include an extra term to the mean, thus:

$$\begin{split} g(\mu_x, \mu_y) + \frac{1}{2} \frac{\partial^2 g(\mu_x, \mu_y)}{\partial \mu_x^2} \mathrm{Var}(\bar{X}) + \frac{1}{2} \frac{\partial^2 g(\mu_x, \mu_y)}{\partial \mu_y^2} \mathrm{Var}(\bar{Y}) \\ + \frac{\partial^2 g(\mu_x, \mu_y)}{\partial \mu_x \partial \mu_y} \mathrm{Cov}(\bar{X}, \bar{Y}) + \dots \end{split}$$

which, in the RIS case, reduces to

$$\begin{split} g(\mu_x, \mu_y) + \frac{1}{2} \frac{\partial^2 g(\mu_x, \mu_y)}{\partial \mu_x^2} \frac{\sigma_x^2}{N} + \frac{1}{2} \frac{\partial^2 g(\mu_x, \mu_y)}{\partial \mu_y^2} \frac{\sigma_y^2}{N} + \\ + \frac{\partial^2 g(\mu_x, \mu_y)}{\partial \mu_x \partial \mu_y} \frac{\mu_{1,1}}{N} + \dots \end{split}$$

 $(\mu_{1,1} \text{ indicates the corresponding } central \text{ moment, namely the covariance between } one X \text{ and the corresponding } Y).$

It is also necessary to incorporate the $\frac{1}{\sqrt{N}}$ proportional skewness, based on the following third central moment of $g(\bar{X})$ (assuming, for simplicity, that there is only one \bar{X} involved):

$$\begin{split} &\mathbb{E}\left[\left(g(\bar{X}) - g(\mu_x) - \frac{1}{2}\frac{\partial^2 g(\mu_x)}{\partial \mu_x^2}\mathrm{Var}(\bar{X}) + \ldots\right)^3\right] \simeq \\ &\left(\frac{\partial g(\mu_x)}{\partial \mu_x}\right)^3\mathbb{E}\left[\left(\bar{X} - \mu_x\right)^3\right] + \frac{3}{2}\left(\frac{\partial g(\mu_x)}{\partial \mu_x}\right)^2\frac{\partial^2 g(\mu_x)}{\partial \mu_x^2}\left(\mathbb{E}\left[\left(\bar{X} - \mu_x\right)^4\right] - \mathrm{Var}(\bar{X})^2\right) + \ldots \end{split}$$

For RIS, this yields

$$\begin{split} & \left(\frac{\partial g(\mu_x)}{\partial \mu_x}\right)^3 \frac{\mu_3}{N^2} + \frac{3}{2} \left(\frac{\partial g(\mu_x)}{\partial \mu_x}\right)^2 \frac{\partial^2 g(\mu_x)}{\partial \mu_x^2} \left(\frac{\mu_4}{N^3} + 3\frac{N(N-1)\sigma^4}{N^4} - \frac{\sigma^4}{N^2}\right) + \dots \\ & \simeq \left(\frac{\partial g(\mu_x)}{\partial \mu_x}\right)^3 \frac{\mu_3}{N^2} + 3 \left(\frac{\partial g(\mu_x)}{\partial \mu_x}\right)^2 \frac{\partial^2 g(\mu_x)}{\partial \mu_x^2} \frac{\sigma^4}{N^2} + \dots \end{split}$$

where μ_3 and μ_4 are the *central* moments of the X distribution.

For the special case of RIS and $g(\bar{X})$, we get the following formulas

$$\mathbb{E}\left[g(\bar{X})\right] \simeq g(\mu) + \frac{g''(\mu)\sigma^2}{2N} + \dots$$

$$\operatorname{Var}\left[g(\bar{X})\right] \simeq \frac{g'(\mu)^2\sigma^2}{N} + \frac{g'(\mu)g''(\mu)\kappa_3 + \frac{1}{2}g''(\mu)^2\sigma^4 + g'(\mu)g'''(\mu)\sigma^4}{N^2} + \dots$$

$$\mathbb{E}\left\{\left[g(\bar{X}) - \mu_g\right]^3\right\} \simeq \frac{g'(\mu)^3 \cdot \mu_3 + 3g'(\mu)^2g''(\mu)\sigma^4}{N^2} + \dots$$

$$\mathbb{E}\left\{\left[g(\bar{X}) - \mu_g\right]^4\right\} \simeq \frac{3g'(\mu)^4\sigma^4}{N^2} + \frac{g'(\mu)^4\kappa_4 + 18g'(\mu)^3g''(\mu)\kappa_3\sigma^2 + 15g'(\mu)^2g''(\mu)^2\sigma^6 + 10g'(\mu)^3g'''(\mu)\sigma^6}{N^3} + \dots$$

where κ_i are cumulants of the X distribution, and μ_g denotes the expected value of $g(\bar{X})$.

The last expression can be easily converted into the corresponding 4^{th} cumulant of $g(\bar{X})$, getting

$$\frac{g'(\mu)^2 \left[g'(\mu)^2 \kappa_4 + 12 g'(\mu) g''(\mu) \kappa_3 \sigma^2 + 12 g''(\mu)^2 \sigma^6 + 4 g'(\mu) g'''(\mu) \sigma^6\right]}{N^3} + \dots$$